

Death and its Determinants

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Declan French (CoE NI) and
Colin O'Hare (QUMS NI)

Outline of presentation

- Motivation and context within modeling mortality
- Review of existing Stochastic atheoretical models
- Data sources and quality
- The methodology
- The fitting and forecasting results
- Annuity pricing implications
- Conclusions from the results and further research

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Motivation and context within modeling mortality

- Longevity risk a significant risk to be managed
- Significant literature identifying and modeling the improvement in mortality
- Stochastic “extrapolative” models well established
- Lee Carter (1992), Brouhns et al (2002), Renshaw and Haberman (2003,2006), Cairns et al (2006,2009), Currie (2006), Girosi and King (2005),....
- Common trends identified and forecast

Motivation continued

- Incorporating observable factors that may explain trends is missing from the literature
- atheoretical models identified evidence for and added factors to improve the fit
- Little attention is paid to the forecasting ability
- Little attention is paid to a rigorous approach to identifying the necessary number of parameters

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Review of existing Stochastic atheoretical models

- Lee Carter identified a single time trend in mortality rates

$$m_{x,t} = \exp(\alpha_x + \beta_x \kappa_t + \varepsilon_t)$$

- Additional trend in cohort identified in Renshaw & Haberman (2003,2006) and Currie (2006)

$$m_{x,t} = \exp(\alpha_x + \beta_x^1 \kappa_t^1 + \beta_x^2 \gamma_{t-x} + \varepsilon_t) \quad m_{x,t} = \exp(\alpha_x + \kappa_t^1 + \gamma_{t-x} + \varepsilon_t)$$

- Cairns (2006) linked the second trend to time and developed a series of models with two time trends and cohort effects

$$\text{logit}(q_{x,t}) = \log(q_x / 1 - q_x) = \kappa_t^1 + \kappa_t^2 (x - \bar{x}) + \varepsilon_t$$

- Plat (2009) combined all trends with 4 factors (3 time related, 1 cohort)

$$m_{x,t} = \exp(\alpha_x + \kappa_t^1 + \kappa_t^2 (\bar{x} - x) + \kappa_t^3 (\bar{x} - x)^+ + \gamma_{t-x} + \varepsilon_t)$$

Girosi and King (2008)

- Extrapolative model in line with previous stochastic models
- Additional information about time and age incorporated via priors
- Bayesian structure to the model:

$$m_{x,t} \square N\left(\mu_{x,t}, \frac{\sigma_x^2}{b_{x,t}}\right) \quad x = 1, \dots, N, t = 1, \dots, T$$

$$\mu_{x,t} = Z_{x,t} \beta_x$$

- Prior knowledge encapsulates: “similar” cross sections should have “similar” coefficients.

$$P(\beta | \theta) \propto \exp\left(-\frac{1}{2} H^\beta [\beta, \theta]\right)$$

$$H^\beta [\beta, \theta] \equiv \frac{1}{2} \sum s_{i,j} \|\beta_i - \beta_j\|_\theta^2$$

King and Soneji (2011)

- Inclusion of exogenous variables in addition to time trends
- Exogenous variables (smoking prevalence) lagged by 25 years taken from the literature
- They conclude that including risk factors (obesity and smoking) improves forecasts and shows mortality rate decline to be greater than that previously predicted.
- R package called “YourCast” used to applied this model
- Implemented in R and available from: <http://gking.harvard.edu/yourcast>

Outline of presentation

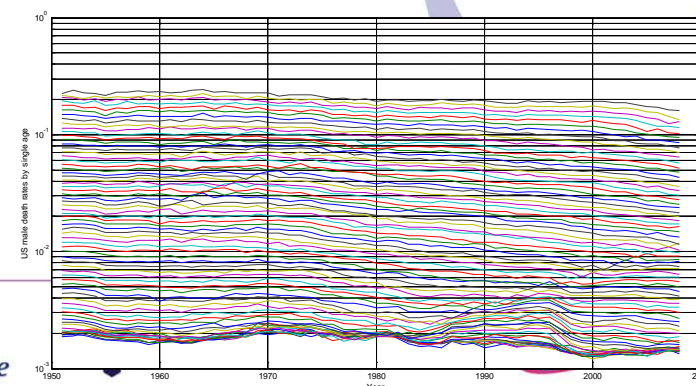
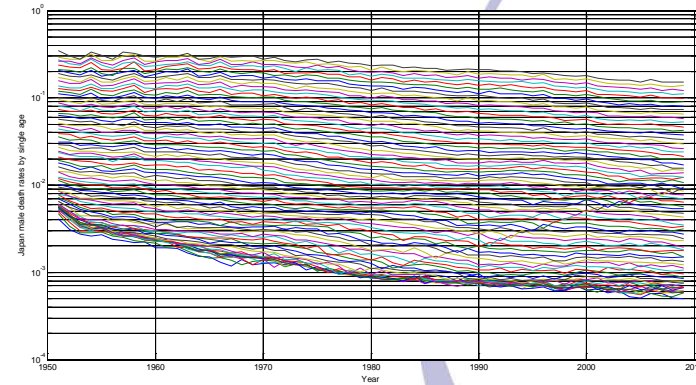
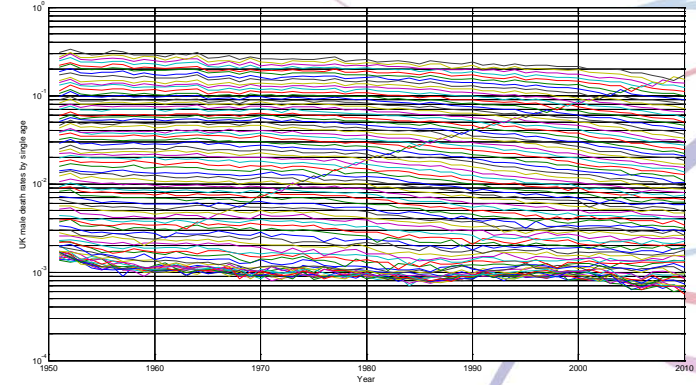
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Data

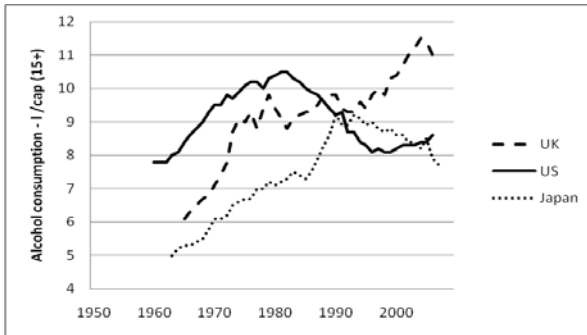
- Mortality data taken from the Human Mortality database
<http://www.mortality.org>
- Data on possible determinants of health taken from OECD health data 2009
<http://www.oecd.org>
- Study includes U.S., U.K. and Japan
 - Common characteristic: all with developed healthcare systems, all non-tropical countries
 - Differing characteristic: culture, diet and the significance of private vs public health expenditure
- Data taken from 1970 – 2000 and forecast from 2001-2006 inclusive and back testing carried out.
- Exogenous variables chosen included: Alcohol consumption, Tobacco consumption, GDP, Health expenditure, total fat intake, total fruit and vegetable consumption.

Data – Mortality rates

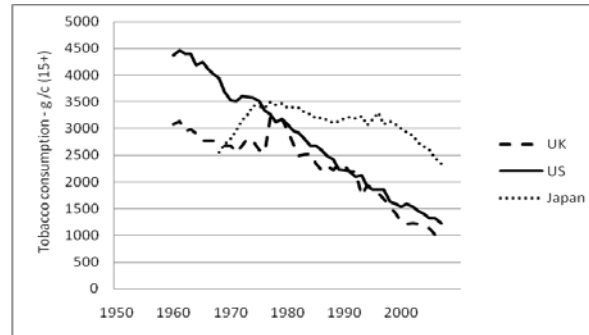
- Log mortality rates $m_{x,t} = \ln\left(\frac{D_{x,t}}{E_{x,t}}\right)$ for the U.K., U.S. and Japan.
- Declining mortality rates
- Cohort effect visible



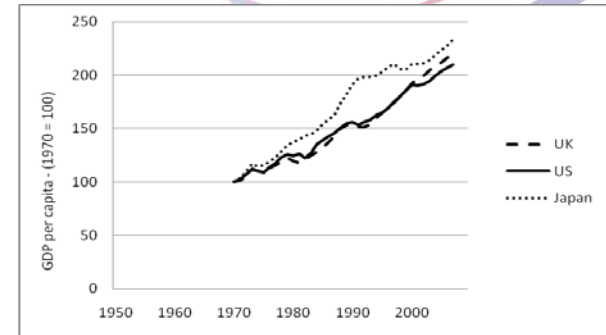
Data – exogenous variables



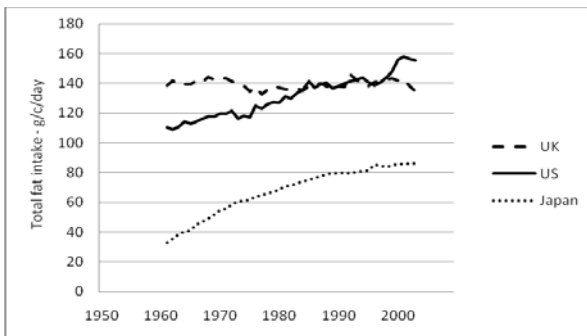
Alcohol



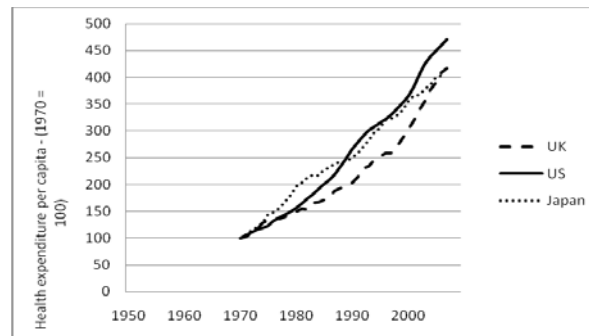
Tobacco



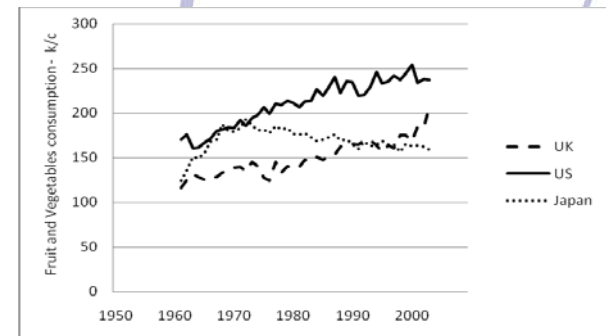
GDP



Fat Intake



Health Expenditure



Fruit and Veg intake

Data – exogenous variables

J	Mean	Standard Deviation	Definition
Alcohol	(UK) 9.3 (US) 9.5 (Japan) 7.8	0.6 0.8 1.0	Annual consumption of pure alcohol in litres, per person, aged 15 years and over
Tobacco	2349 2645 3227	516 667 147	Annual consumption of tobacco items (eg cigarettes, cigars) in grams per person aged 15 years and over
Fat	138.9 133.5 73.1	3.0 10.0 8.9	Total fat (grams per capita per day)
Fruit & Veg	151.8 219.7 173.1	14.3 17.6 8.8	All fruit and vegetable consumption (except wine) in kilos per capita
GDP	11896 25420 2,994,819	2275 4712 707,685	Gross domestic product per capita in national currency units at 2000 price levels
Health exp	712 2787 191,957	222 1098 63,732	Total health expenditure (private and public) per capita in national currency units at 2000 price levels

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Overall methodology

- Step 1 – factor structure analysis
- Step 2 – linking of latent factors with exogenous factors
- Step 3 – forecasting of exogenous variables
- Step 4 – King and Soneji (2011) modeling with forecast exogenous variables

Factor structure analysis

- Principal components method applied
- Stopping rule applied to limit the number of extracted factors
- Method of Bai and Ng (2002) applied to determine the number of factors
- Information criterion used to limit factor extraction
- The number of factors “ r ” which minimises this gives the estimated number of factors.

Factor structure analysis

- Describe the N variables with a smaller number of factors $\mathbf{F} = (F_1, \dots, F_r)$

$$\text{Data } x_{it} = \lambda_i \mathbf{F} + e_{it} \quad i=1, \dots, N, t=1, \dots, T$$

- Using PC, factor estimates are given by the first r eigenvectors of the matrix $XX'/(NT)$
- Information criterion used to limit factor extraction

$$IC_p(r) = \log \hat{\sigma}^2(r) + r \frac{N+T}{NT} \ln[\min(N, T)] \quad \text{where } \hat{\sigma}^2(r) = \frac{1}{NT} \sum_{i=1}^N \sum_{t=1}^T \tilde{e}_{it}^2$$

- The number of factors “ r ” which minimises this gives the estimated number of factors.

Linking to the observed variables

- Matrix G_t of observed variables be expressed as a linear combination of the r latent variables F_t ?
- We may be able to express each variable as an exact combination or approximate combination

$$G_{jt} = \delta_j' F_t \quad \forall t$$

$$G_{jt} = \delta_j' F_t + \varepsilon_{jt} \quad \forall t$$

- Let $\hat{\delta}$ be a least squares estimate of δ and calculate $\hat{G}_{jt} = \hat{\delta}_j' \tilde{F}_t \quad \forall t$

$$\tau_t(j) = \frac{\hat{G}_{jt} - G_{jt}}{\left(\text{var } \hat{G}_{jt}\right)^{1/2}}$$

Linking – Exact case

- Test the exogenous variables individually or as a group
- Test of individual exogenous variable consists of test statistic
(1) a test of the proportion of the time series for which the linear combination estimate deviates from the exogenous variable

$$A(j) = \frac{1}{T} \sum_1^T 1(\hat{\tau}_t(j) > \phi_\alpha)$$

(2) a test of how far the estimate is from the exogenous variable over the whole time series

$$M(j) = \text{Max}_{1 \leq t \leq T} |\hat{\tau}_t(j)|$$

Linking – approximate case

- In the approximate case we are saying that there is some noise in the linear relationship between the exogenous variable and the latent factors.
- The two equivalent tests in this case are the noise to signal ratio and the coefficient of determination.

$$NS(j) = \frac{\text{var}(\hat{\mathcal{E}}(j))}{\text{var}(\hat{G}(j))}$$

$$R^2(j) = \frac{\text{var}(\hat{G}(j))}{\text{var}(G(j))}$$

- In testing the group of exogenous variables we first look for the largest canonical relationship between linear combinations of exogenous variables. We then repeat the process but looking for a second relationship uncorrelated to the first etc.

Linking – Canonical Correlations

- Testing the group as a set, the canonical correlations between G and F are considered.
- The first canonical correlation, ρ_1 , is the largest correlation that can be found for linear combinations of G and F.
- The second canonical correlation, ρ_2 , is the largest correlation that can be found from linear combinations of G and F uncorrelated with those giving the first canonical correlation, and so on.
- If all the m variables in are exact factors then the canonical correlations will all be unity. If the m variables are linearly dependent then the number of non-zero canonical correlations will be less than m.

Forecasting exogenous determinants

- Box-Jenkins methodology used to forecast exogenous variables
- Moving average terms ignored as these give poorer forecasts
- Schwarz Information Criteria used to determine the number of autoregressive components
- GDP and Health expenditure differenced first before applying the SIC

King and Soneji (2011) – The Model

- The King and Soneji model will include the exogenous variables selected

$$m_{x,t} \square N \left(\mu_{x,t}, \frac{\sigma_x^2}{b_{x,t}} \right) \quad x = 1, \dots, N, t = 1, \dots, T$$

$$\mu_{x,t} = Z_{x,t} \beta_x$$

- Exogenous variables built into the model through the covariate beta vector, for example:

$$m_{x,t} = \beta_x^0 + \beta_x^1 year_t + \left(\beta_x^{gdp} gdp + \beta_x^{alc} alc + \dots \right) + \varepsilon_{x,t}$$

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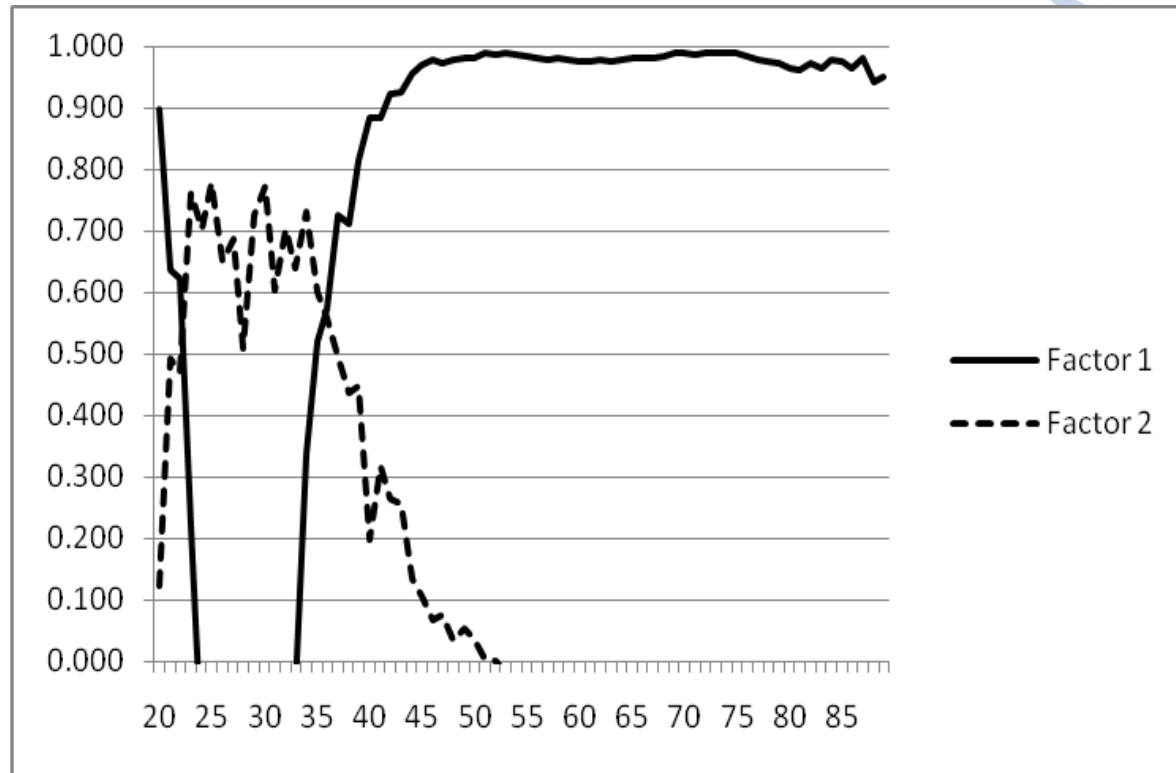
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Latent factor structure

- Using the methods of Bai and Ng (2002) with appropriate stopping rules we find the number of factors necessary to explain the data is:
 - 2 factors explains 86% of variability for U.K.
 - 4 factors explains 98% of the variability for U.S.
 - 4 factors explains 98% of the variability for Japan

Latent factor structure

U.K. (2 factors) – factor 2 explaining the younger ages, factor 1 explaining the older ages



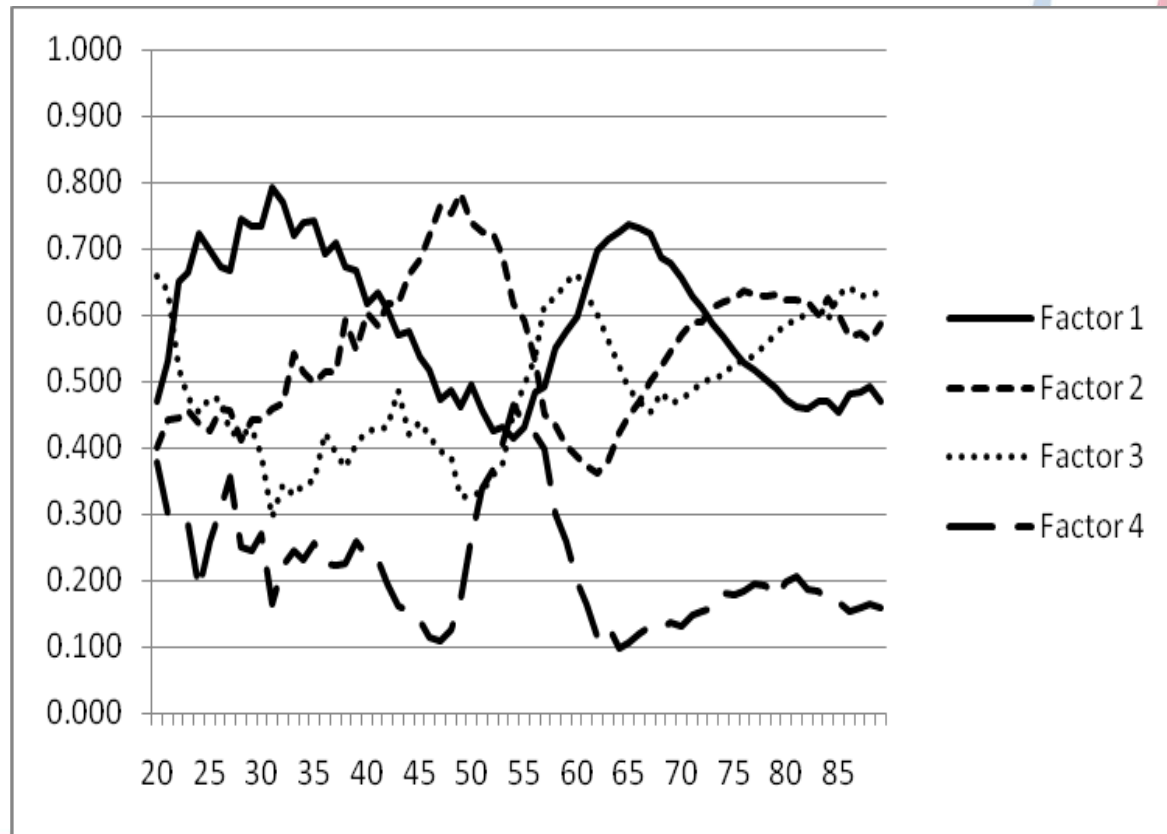
Latent factor structure

- U.S. (4 factors) - Factors 1 and 4 explaining older ages, factors 2 and 3 explaining younger ages



Latent factor structure

- Japan (4 factors) – no particular link between each factor and the age to which it may apply.



Canonical Linking results

U.K.

j	A(j)	M(j)	R ² (j)	NS(j)	(k) ²
Alcohol	0.839	49.14	0.546 (0.310, 0.782)	0.832	0.991 (0.987, 0.997)
Tobacco	0.871	24.62	0.809 (0.688, 0.930)	0.236	0.323 (0.052, 0.594)
Fat	0.645	26.43	0.313 (0.042, 0.584)	2.195	-
Fruit & Veg	0.871	28.02	0.815 (0.698, 0.933)	0.227	-
GDP	0.645	10.43	0.967 (0.944, 0.990)	0.035	-
Health exp	0.484	14.80	0.970 (0.949, 0.991)	0.031	-

U.S.

j	A(j)	M(j)	R ² (j)	NS(j)	(k) ²
Alcohol	0.323	5.94	0.976 (0.960, 0.993)	0.024	0.999 (0.998, 1.000)
Tobacco	0.710	7.53	0.991 (0.984, 0.997)	0.009	0.951 (0.918, 0.985)
Fat	0.806	25.82	0.911 (0.850, 0.971)	0.098	0.366 (0.096, 0.636)
Fruit & Veg	0.935	44.99	0.878 (0.798, 0.959)	0.139	0.158 (0.001, 0.352)
GDP	0.806	14.56	0.975 (0.958, 0.992)	0.025	-
Health exp	0.419	5.20	0.997 (0.994, 0.999)	0.003	-

Japan

j	A(j)	M(j)	R ² (j)	NS(j)	(k) ²
Alcohol	0.484	9.494	0.978 (0.962, 0.993)	0.023	0.995 (0.992, 0.999)
Tobacco	0.484	6.893	0.804 (0.680, 0.928)	0.244	0.938 (0.895, 0.980)
Fat	0.581	8.961	0.992 (0.986, 0.998)	0.008	0.786 (0.652, 0.919)
Fruit & Veg	0.839	39.868	0.795 (0.666, 0.924)	0.258	0.310 (0.039, 0.580)
GDP	0.484	8.338	0.993 (0.988, 0.998)	0.007	-
Health exp	0.548	7.130	0.992 (0.987, 0.998)	0.008	-

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ARIMA Modeling of exogenous variables

No allowance for Moving Average terms in the ARIMA forecast
Used SIC to test for various Autoregressive terms

U.K.

GDP – ARIMA(2,1,0)

U.S.

Health – ARIMA(1,1,0)

Alcohol – ARIMA(3,0,0)

Japan.

Health – ARIMA(0,1,0)

Fat Intake – ARIMA(1,0,0)

Alcohol – ARIMA(1,0,0)

Measures of quality

- The average error E1 – this equals the average of the standardized errors, i.e. Error_x/n , where n = the number of ages included in the summation, that is the mean of the differences. This is a measure of the overall bias in the projections.
- The average absolute error E2 – this equals the average of $|\text{Error}_x|$, that is the mean of the absolute differences. This is a measure of the magnitude of the differences between the actual and projected rates.
- The standard deviation of the error E3 – this equals the square root of the average of the squared errors (Error_x^2), the root mean squares of the differences between the projected and actual rates.

Forecasting results

U.K.

Model	E1	E2	E3
Lee Carter (1992)	10.600%	10.652%	13.920%
Girosi and King (2008)	10.349%	10.362%	12.413%
King and Soneji (2011)	9.254%	9.254%	11.551%

U.S.

Model	E1	E2	E3
Lee Carter (1992)	8.268%	8.325%	11.161%
Girosi and King (2008)	7.462%	7.700%	10.219%
King and Soneji (2011)	6.661%	7.016%	8.841%

Japan

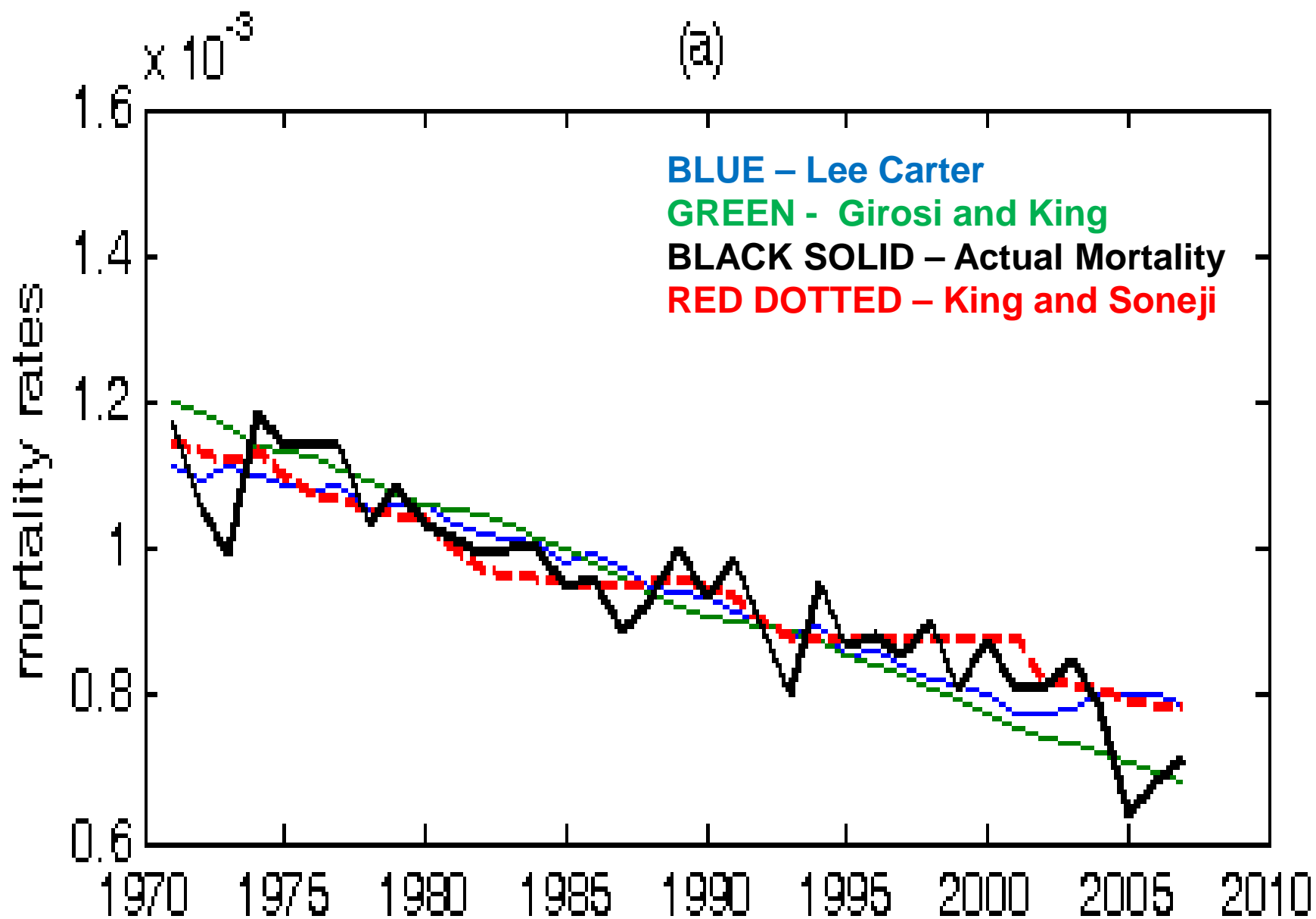
Model	E1	E2	E3
Lee Carter (1992)	8.579%	8.579%	10.256%
Girosi and King (2008)	8.790%	8.793%	11.501%
King and Soneji (2011)	4.729%	4.729%	6.007%

Fitting results

Model	E1	E2	E3
U.K. Lee Carter (1992)	0.247%	3.501%	4.900%
Girosi and King (2008)	0.138%	3.903%	5.134%
King and Soneji (2011)	0.117%	3.757%	4.940%

Model	E1	E2	E3
U.S. Lee Carter (1992)	0.333%	3.832%	5.784%
Girosi and King (2008)	0.040%	4.119%	5.986%
King and Soneji (2011)	-0.089%	2.551%	3.402%

Model	E1	E2	E3
Japan Lee Carter (1992)	0.394%	3.857%	5.051%
Girosi and King (2008)	0.238%	4.994%	6.461%
King and Soneji (2011)	0.072%	2.939%	3.853%



U.K. results

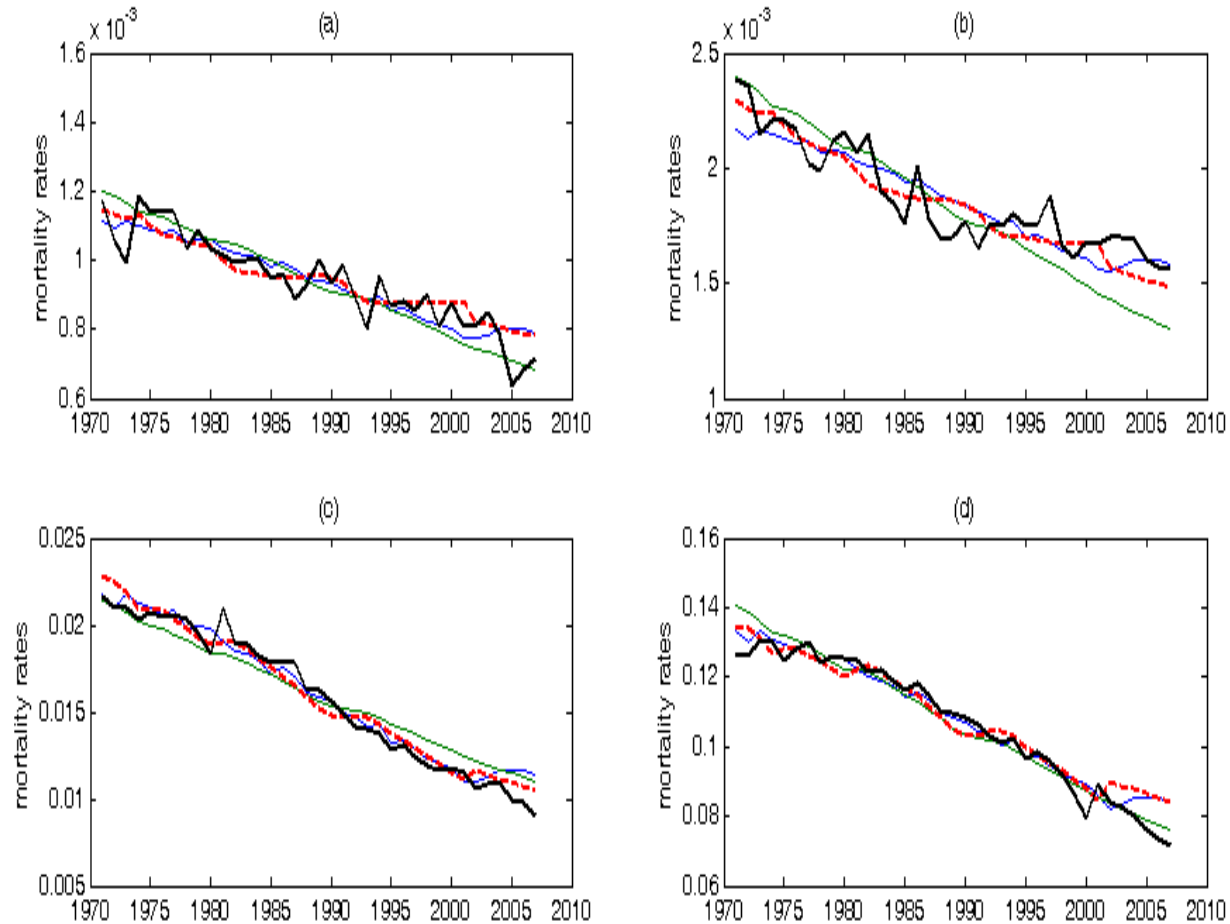


Figure 14 U.K. mortality rates fitted between 1970–2000, and forecast from 2001 - 2006 for the Lee Carter, (blue), Giosi and King (green), King and Soneji, (bold red dashed) models and actual mortality rates 1970–2006 (bold black) for males aged (a) 20, (b) 40, (c) 60 and (d) 80

U.S. results

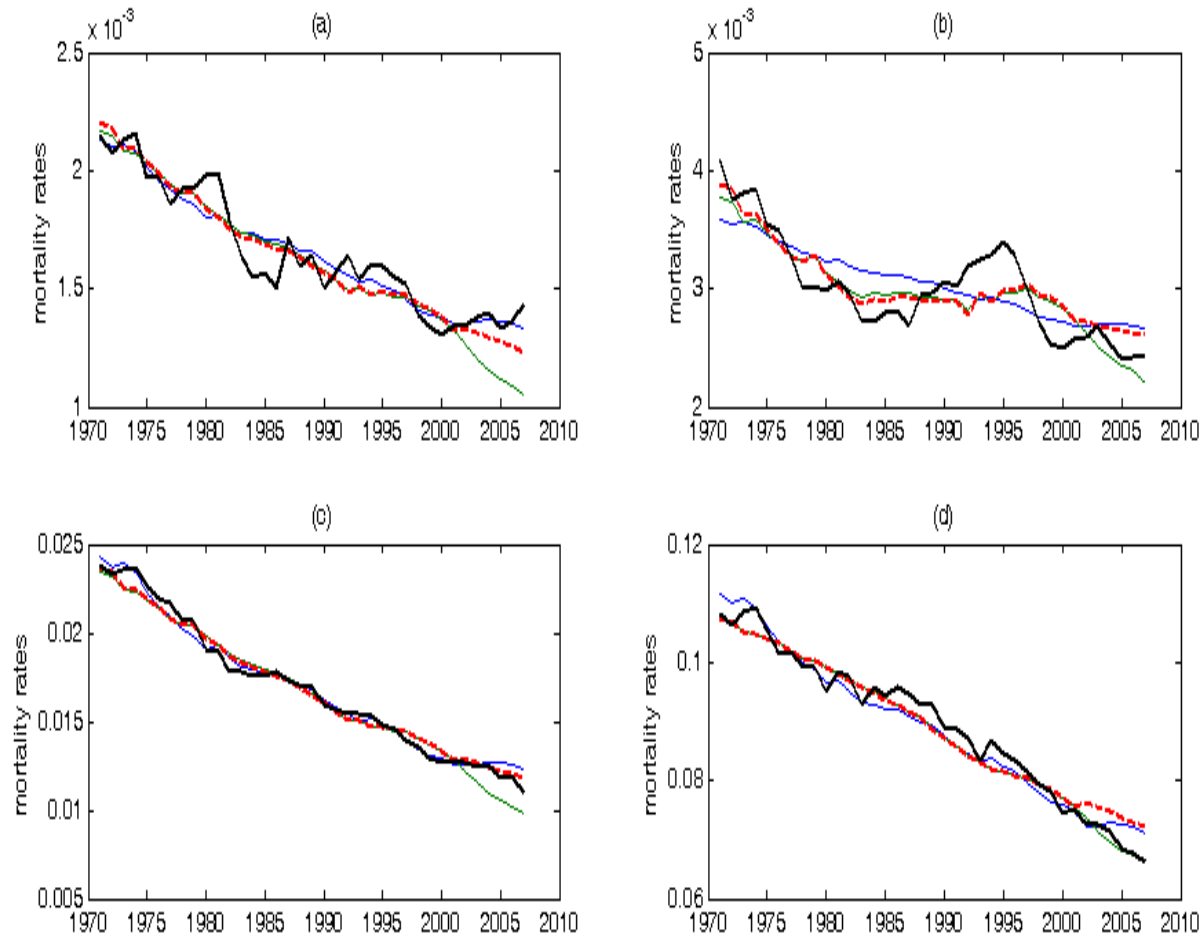


Figure 15 U.S mortality rates fitted between 1970–2000, and forecast from 2001 - 2006 for the Lee Carter, (blue), Girosi and King (green), King and Soneji, (bold red dashed) models and actual mortality rates 1970–2006 (bold black) for males aged (a) 20, (b) 40, (c) 60 and (d) 80

Japan results

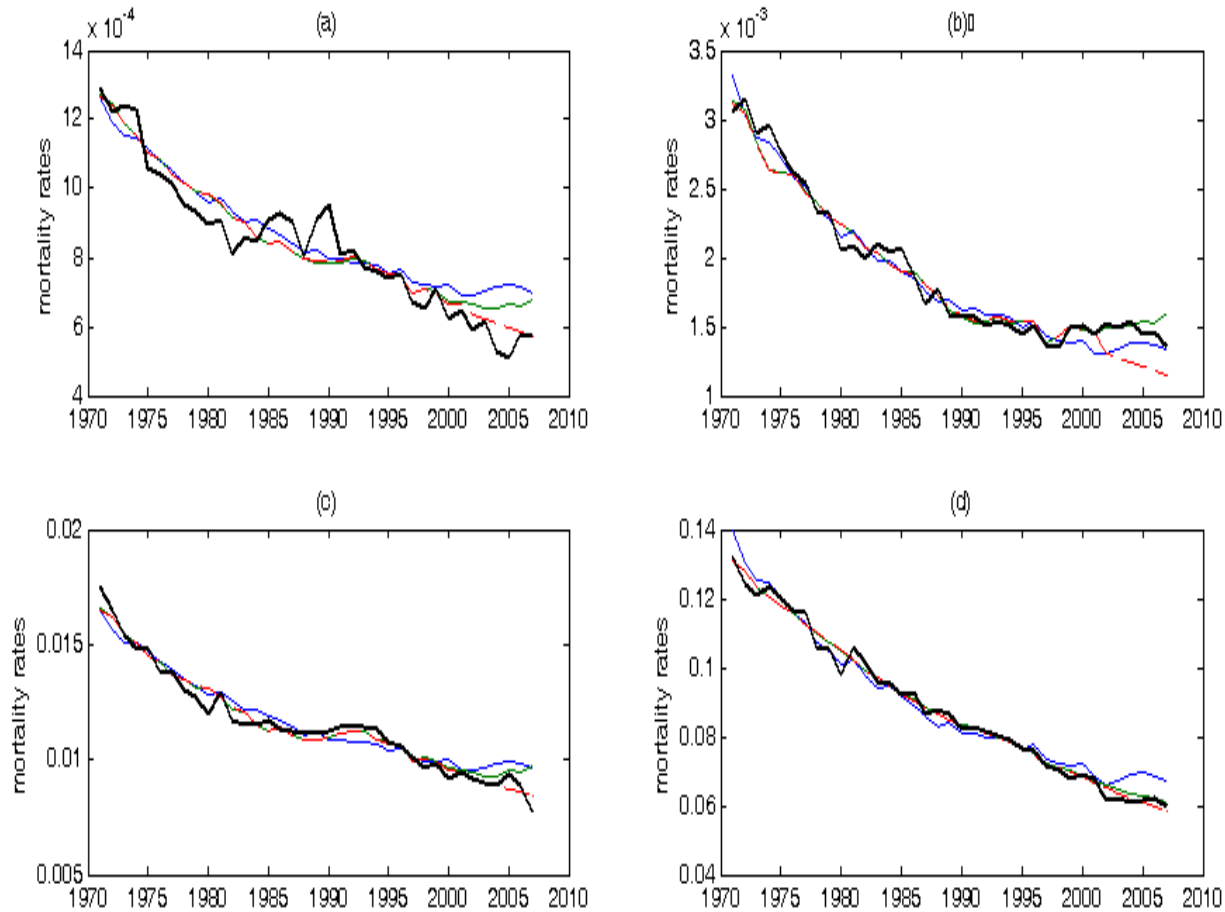
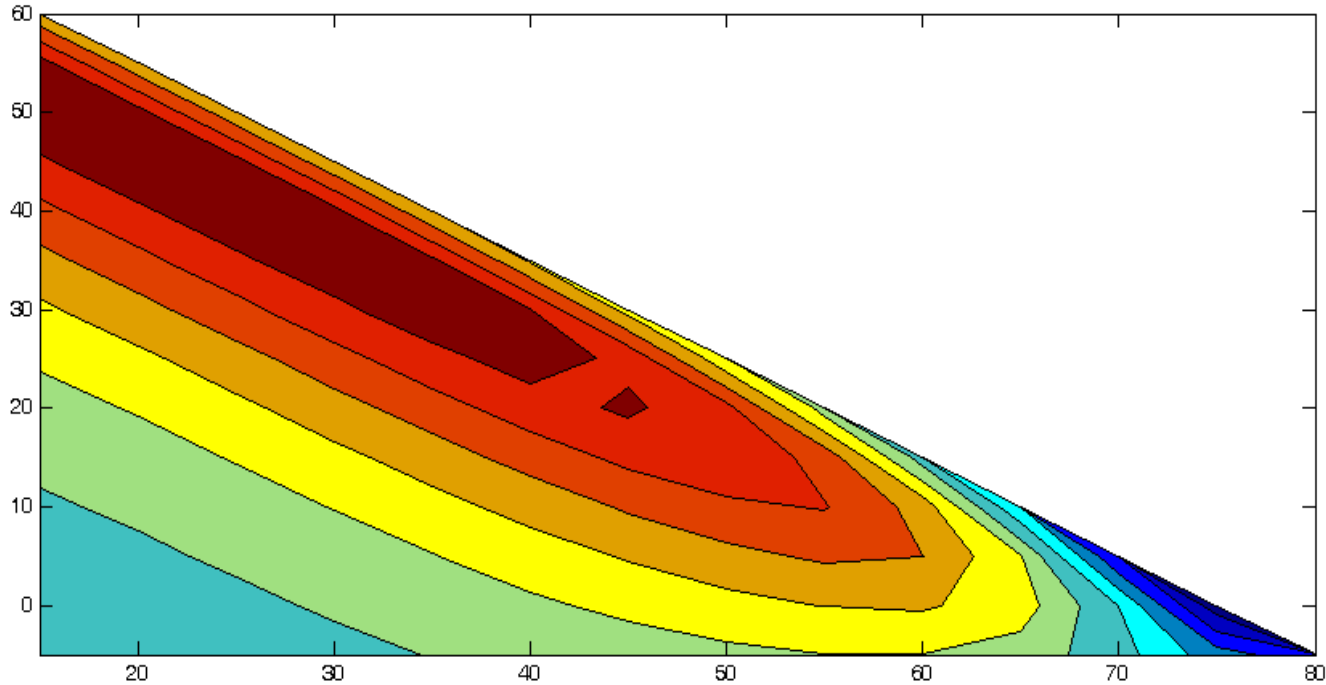


Figure 16 Japan mortality rates fitted between 1970–2000, and forecast from 2001 - 2006 for the Lee Carter, (blue), Girosi and King (green), King and Soneji, (bold red dashed) models and actual mortality rates 1970–2006 (bold black) for males aged (a) 20, (b) 40, (c) 60 and (d) 80

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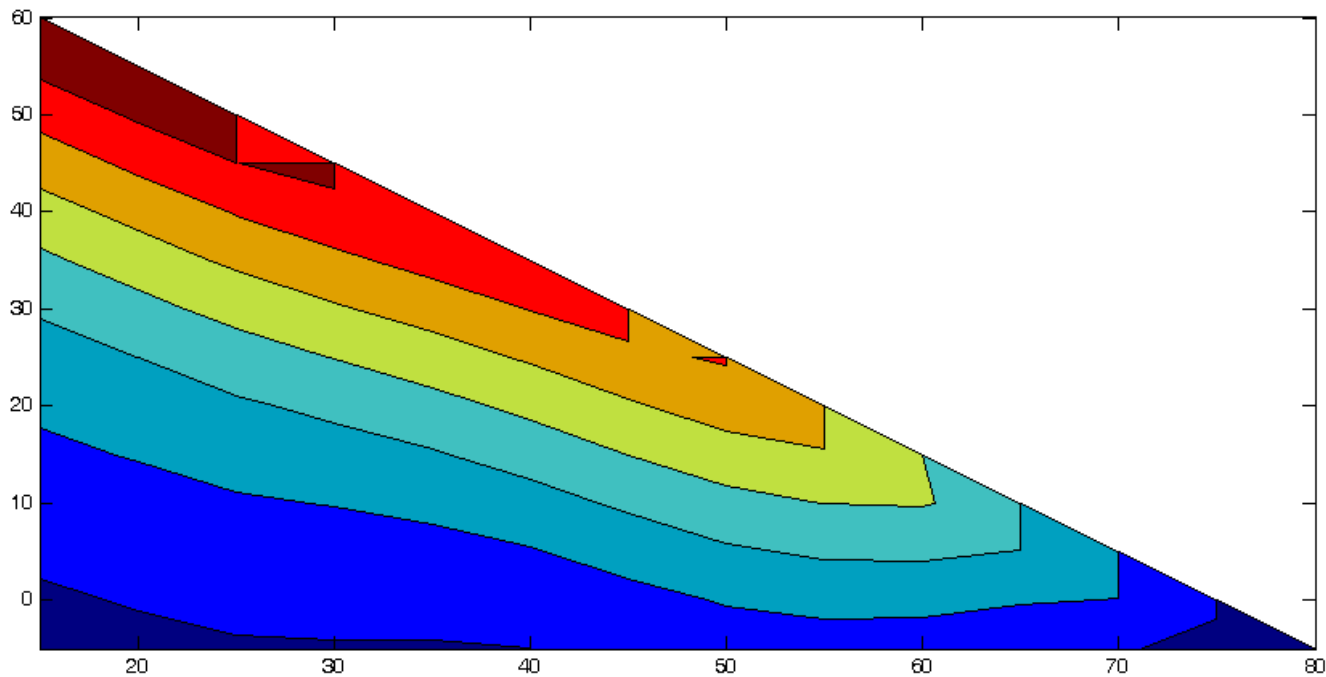
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Annuity pricing implications – U.K.



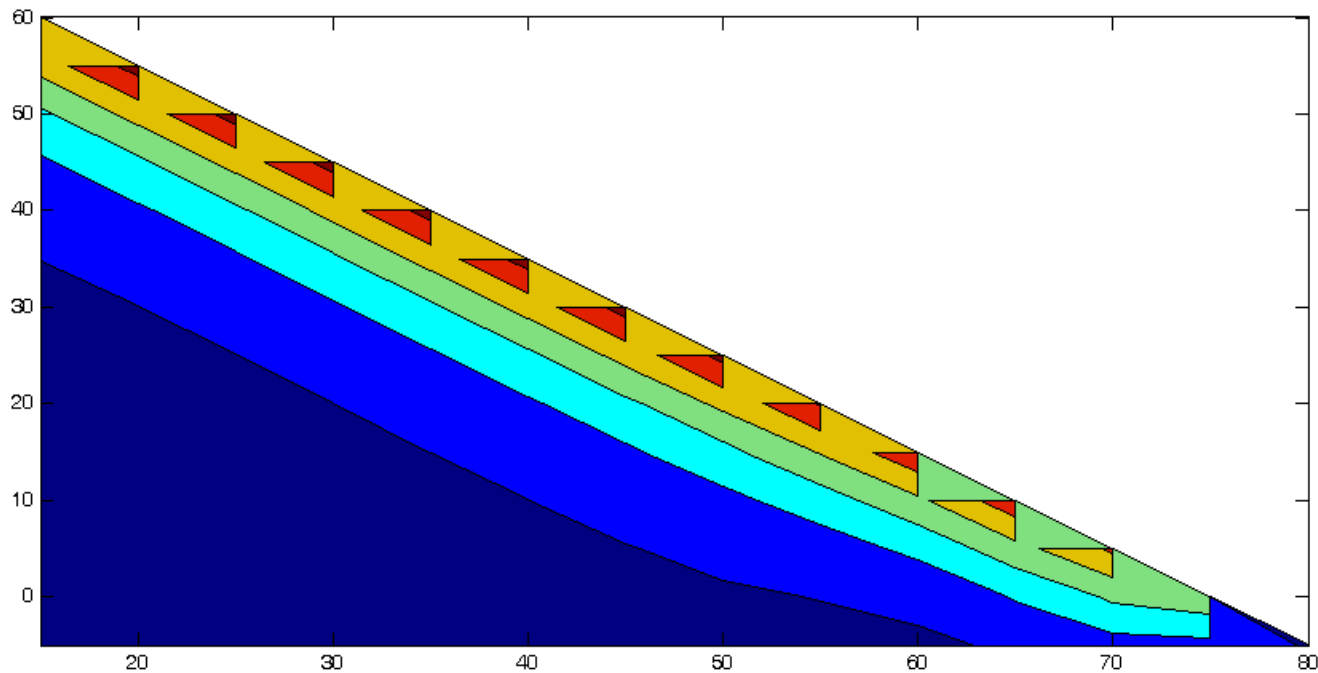
Annuity pricing implications – U.S.

- King and Soneji(2011) vs Lee Carter(1992)



Annuity pricing implications – Japan

- King and Soneji(2011) vs Lee Carter(1992)



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Conclusions from the results

- Our model forecasts better than atheoretical models
- Exogenous variables can be related to the latent factor structure
- Different data sets can be explained more or less with exogenous variables
- U.K. 1 exogenous variable, U.S. 2 exogenous variables, Japan 3 exogenous variables
- With better data we would hope to improve this.

Conclusions from the results

- Forecasts dramatically improved using exogenous variables. Where we have more exogenous input improvements are much greater – e.g. Japan
- Several cases of 100% bias suggesting that logarithmically transforming the data is not sufficient to linearize the data.
- Impact on Annuity pricing is most severe at older ages and where the deferral period is longer.
- With more exogenous variables there is more of an impact on annuity price.

Further research

- Expansion of the set of exogenous variables as and when data becomes available
- Inclusion of lagged variables where appropriate
- Analysis of longer forecasting periods